

Online Factorization and Partition of Complex Networks by Random Walk

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Background

- Large complex Networks are hard to understand and often not fully/explicitly observable

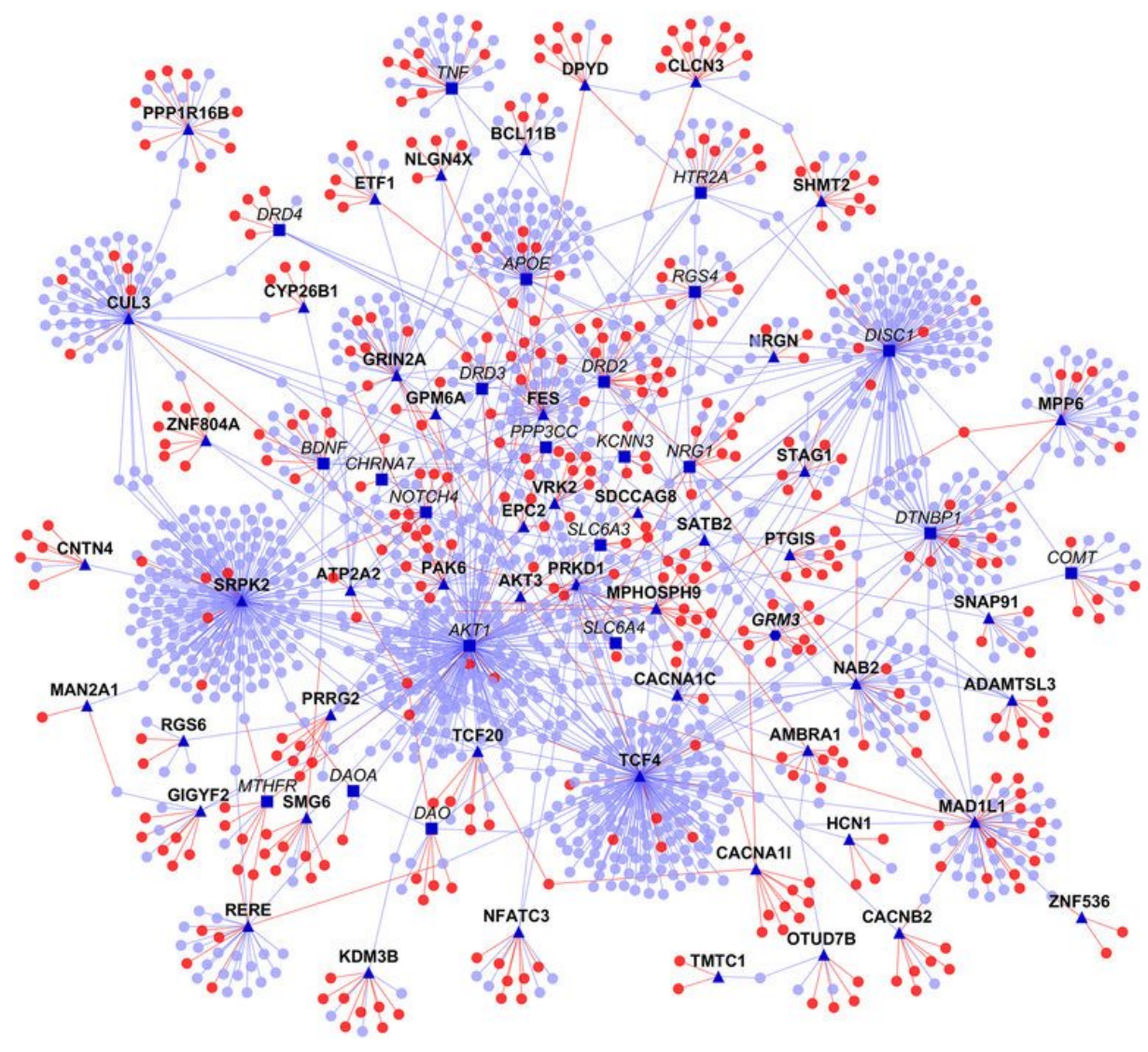
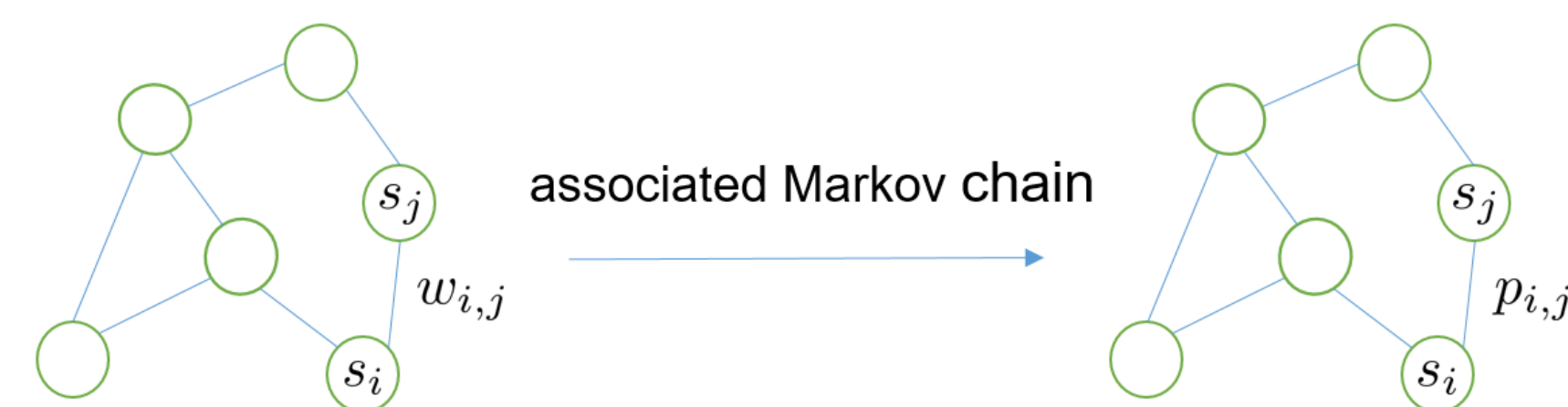


Figure 1: Protein-Protein Interaction Network

- Learning from a "random walk" associated with the network



Network $G = (S, E)$
 Undirected: $w_{i,j} = w_{j,i} \geq 0$
 Connected: $w_i = \sum_j w_{i,j} > 0$

Markov chain $M = (S, p)$
 Transition Probability $p_{i,j} = \frac{w_{i,j}}{\sum_j w_{i,j}}$
 Stationary Distribution $\mu_i = \frac{w_i}{\sum_j w_j}$

Problem of Interest: Given streaming sample trajectory of state transition of the unknown Markov chain:

- how to extract its reduced-order information?
 - how to recover the latent network partition?
- in an **online** fashion with **low** computational and storage cost, and **small** sample complexity.

Model Assumptions

Nearly Low-Rankness: Denote $D = \text{diag}\{\mu_i\}$ and $P = (p_{i,j})$ as stationary distribution and transition matrix, the product DP has a strictly positive singular gap, i.e., it has the decomposition:

$$DP = U \text{diag}(\sigma_1, \dots, \sigma_r) V^T + F \text{ and } \sigma_r > \|F\|_2.$$

Meta-State Structure: The associated Markov chain has a lumpable structure, i.e., there exists a state partition $S = S_1 \cup S_2 \dots \cup S_r$, such that

$$\forall s_k, s_h \in S_i, \forall j \in \{1, \dots, r\}, \sum_{\ell \in S_j} p_{k,\ell} = \sum_{\ell \in S_j} p_{h,\ell}$$

Remark (Weinan et al., 2008). *If Markov chain is lumpable and low rank, then one can recover its exact partition by clustering the reduced-order representations $M = D^{-1}V$.*

Learning Reduced-Order State Representation

Down sampling to handle data dependency:

$$s^{(1)}, s^{(2)}, \dots, s^{(\tau)}, s^{(\tau+1)}, \dots, s^{(2\tau)}, \dots, s^{((b-1)\tau+1)}, \dots, s^{(b\tau)}$$

Construct sample matrix $Z^{(k)} = \mathbb{I}\{s^{(k\tau-1)}, s^{(k\tau)}\}$

- $Z^{(k)}$'s are close to i.i.d. samples and
- $\mathbb{E}[Z^{(k)} | s^{(0)}] \approx DP = U \text{diag}(\sigma_1, \dots, \sigma_r) V^T + F$

Estimating V reduces back to streaming PCA problem!

$$(U^*, V^*) = \arg \max_{\tilde{U}, \tilde{V} \in \mathbb{R}^{m \times r}} \text{tr} [\tilde{U}^T \mathbb{E} Z \tilde{V}]$$

$$\text{subject to } \tilde{U}^T \tilde{U} = \tilde{V}^T \tilde{V} = I_r.$$

Recast into a symmetric problem:

$$W^* = \arg \max_{W \in \mathbb{R}^{2m \times r}} \text{tr} [W^T \mathbb{E} A W] \text{ subject to } W^T W = I_r,$$

$$\text{where } \mathbb{E} A = \begin{bmatrix} 0_{m \times m} & \mathbb{E} Z \\ \mathbb{E} Z^T & 0_{m \times m} \end{bmatrix} \text{ and } W = \frac{1}{\sqrt{2}} [U^T, V^T]^T.$$

Generalized Hebbian Algorithm (GHA)

Input: sample trajectory $s^{(k)}$, block size τ , step size η .

Init: set random $W^{(0)}$ with orthonormal columns; set $k \leftarrow 1$.

Repeat:

Construct matrix $A^{(k)}$ from sample $s^{(k\tau-1)}, s^{(k\tau)}$;
 $W^{(k+1)} \leftarrow W^{(k)} + \eta(A^{(k)} W^{(k)} - W^{(k)} W^{(k)T} A^{(k)} W^{(k)})$
 $s \leftarrow s + 1$;

Until stopping condition is satisfied

Output: $[\hat{U}; \hat{V}] \leftarrow \sqrt{2} W^{(k)}$.

★ **Streaming** algorithm with **low** computational cost $\mathcal{O}(|S|r)$ and **low** storage cost $\mathcal{O}(|S|r)$!

Recovering Network Partition

(1) Run GHA on Markov transition data to obtain $[\hat{U}; \hat{V}]$.

(2) Let $\hat{\mu}$ be the empirical estimate of the stationary distribution, i.e.,

$$\hat{\mu}_i = \sum_{k=1}^n \mathbb{I}(s^{(k)} = i) / n.$$

Let $\hat{D} = \text{diag}(\hat{\mu}_1, \hat{\mu}_2, \dots, \hat{\mu}_m)$. Each row of $\hat{M} = \hat{D}^{-1} \hat{V}$ gives an approximate r -dimensional representation for the corresponding state/vertex.

(3) Find a set of centers $C = \{c_1, c_2, \dots, c_r\} \subset \mathbb{R}^r$ by solving the following problem:

$$\hat{C} = \arg \min_C \sum_{i=1}^m \min_{c \in C} d^2(\hat{M}_{i*}, c),$$

where $d(\hat{M}_{i*}, c_j) = \|\hat{M}_{i*} - c_j\|_2$ is the Euclidean distance.

(4) Partition by assigning each state to its closest center.

Theory

★ **Down sampling size:** By choosing block length τ :

$$\tau \geq \left\lceil \frac{2}{\Phi^2} \log \left(\sqrt{\frac{\mu_{\max}}{\mu_{\min}}} \frac{1}{\eta} \right) \right\rceil,$$

the data samples are sufficiently close to i.i.d. samples drawn from the stationary distribution of the Markov chain.

★ **Convergence analysis for GHA:**

Principle Angle: Given two matrices $U \in \mathbb{R}^{m \times r}$ and $V \in \mathbb{R}^{m \times r}$ with orthonormal columns, the principle angle between U and V is defined as:

$$\Theta(U, V) = \text{diag} [\cos^{-1}(\sigma_1(U^T V)), \dots, \cos^{-1}(\sigma_r(U^T V))]$$

• **ODE Characterization: Global convergence!**

$$\text{Discrete: } \frac{\gamma_{i,t+\eta}^2 - \gamma_{i,t}^2}{\eta} = \mathcal{F}_{i,t} \gamma_{i,t}^2 + O(\eta).$$

$$\text{weakly } \Downarrow \eta \rightarrow 0$$

$$\text{Continuous: } d\gamma_i^2 = b_i \gamma_i^2 dt$$

As $\eta \rightarrow 0$ and $t \rightarrow \infty$,

$$\|\sin \Theta(V, \hat{V}(t))\|_F^2 + \|\sin \Theta(U, \hat{U}(t))\|_F^2 \rightarrow 0.$$

• **SDE Characterization: Convergence rate!**

$$\text{Discrete: } \frac{\zeta_{ij,t+1} - \zeta_{ij,t}}{\sqrt{\eta}} = \mathcal{F}_{ij,t} \zeta_{ij,t} + O(\eta).$$

$$\text{weakly } \Downarrow \eta \rightarrow 0$$

$$\text{Continuous: } d\zeta_{ij} = K_{ij} \zeta_{ij} dt + G_{ij} dB_t,$$

For sufficiently small $\epsilon > 0$, let

$$N = \tilde{\mathcal{O}} \left(\frac{r}{\epsilon(\sigma_r(DP) - \sigma_{r+1}(DP))^2} \right) \text{ and } t = N\eta,$$

we have

$$\lim_{\epsilon \rightarrow 0} \mathbb{P} [\|\sin \Theta(\hat{U}(t), U)\|_F^2 + \|\sin \Theta(\hat{V}(t), V)\|_F^2 > \epsilon] \leq \frac{1}{10}$$

★ **Recovery of network partition:** Suppose that the estimate \hat{U} , \hat{V} , and empirical distribution $\hat{\mu}$ satisfy

$$\|\sin \Theta(\hat{U}, U)\|_F^2 + \|\sin \Theta(\hat{V}, V)\|_F^2 \leq \epsilon \text{ and}$$

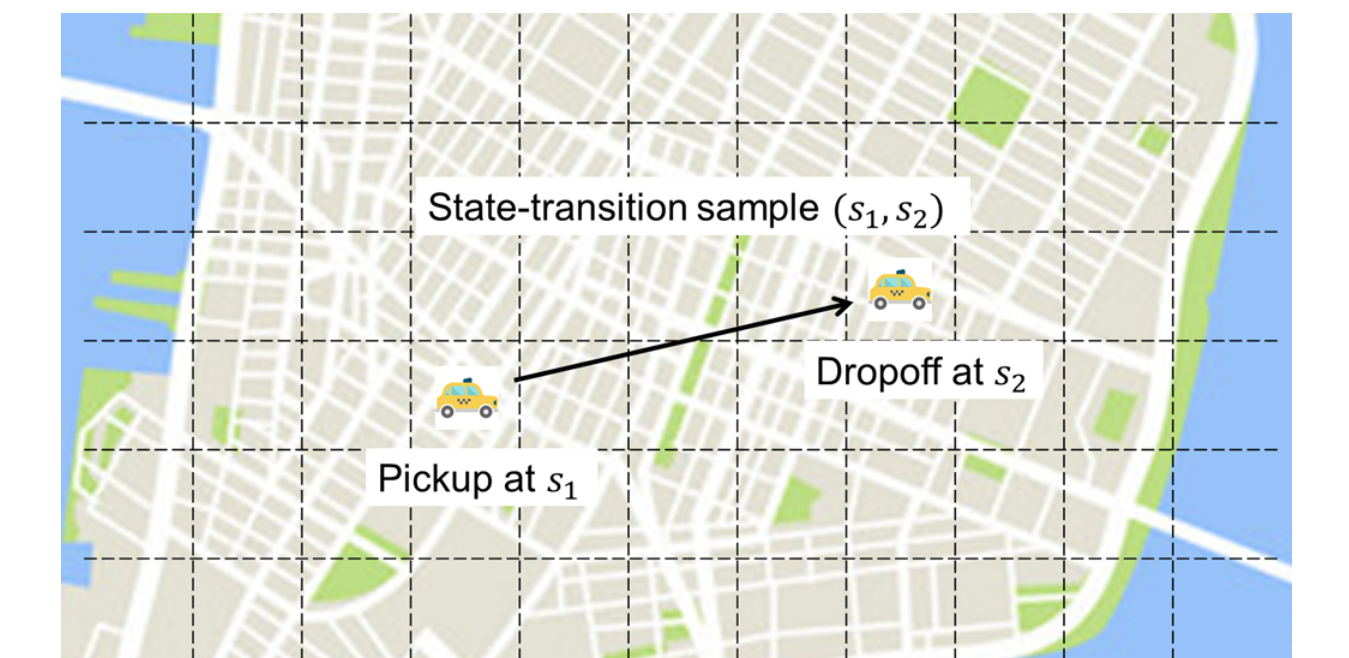
$$\max_{i \in [m]} |\hat{\mu}_i - \mu_i| \leq \sqrt{\epsilon} \mu_i.$$

for some $\epsilon \in (0, 1)$. Let $\hat{M} := \text{diag}(\hat{\mu})^{-1} \hat{V}$ and $M = D^{-1}V$. Then for any $s_i, s_j \in S$,

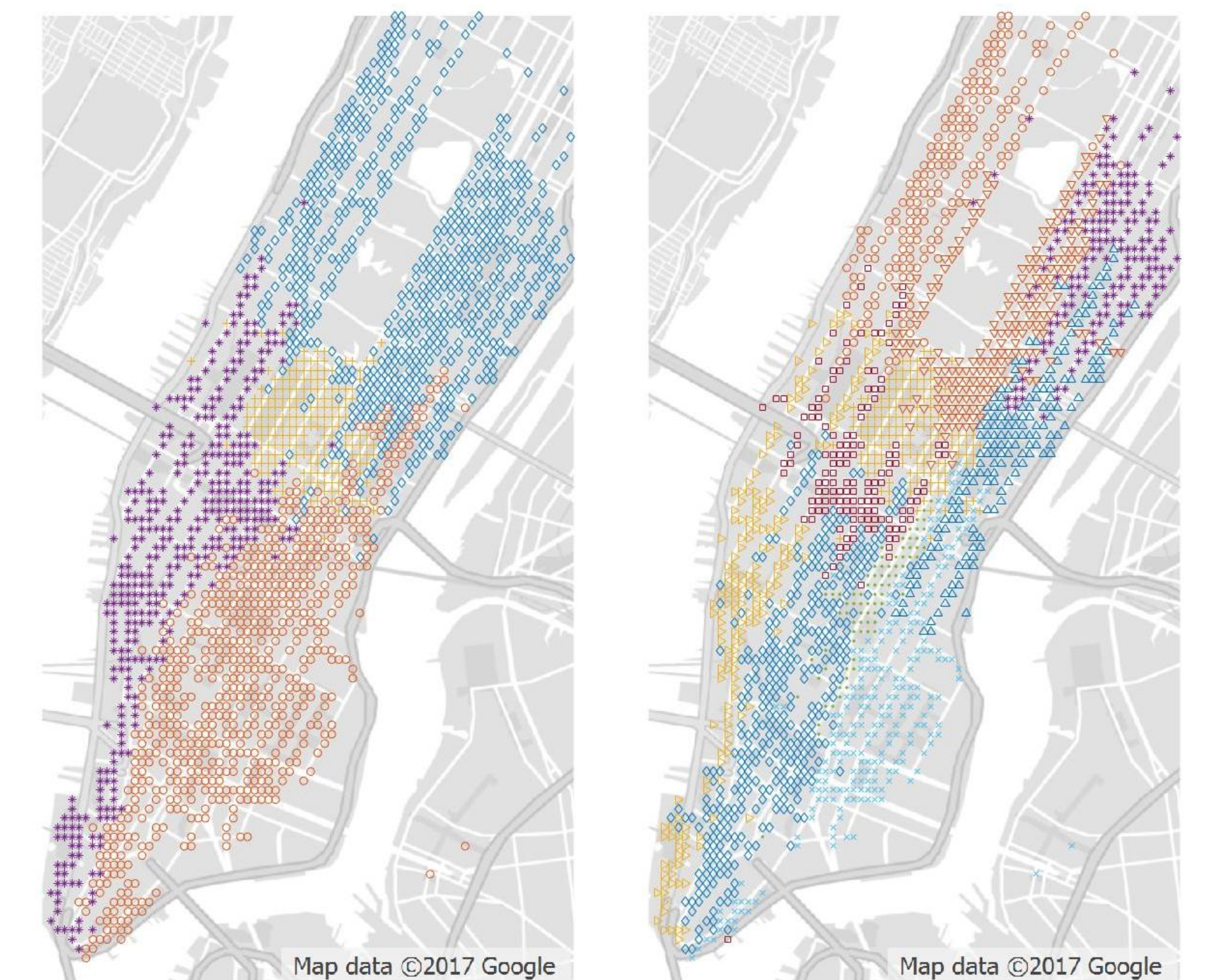
$$\left\| \hat{M}_{s_i*} - \hat{M}_{s_j*} \right\|_2^2 - \left\| M_{s_i*} - M_{s_j*} \right\|_2^2 \leq \frac{C\epsilon}{\mu_{\min}^2}.$$

Experiments

We look at Manhattan taxi data with 1.1×10^7 trip records of NYC Yellow cabs from January 2016. Each entry records the coordinates of the pick-up and drop-off locations.

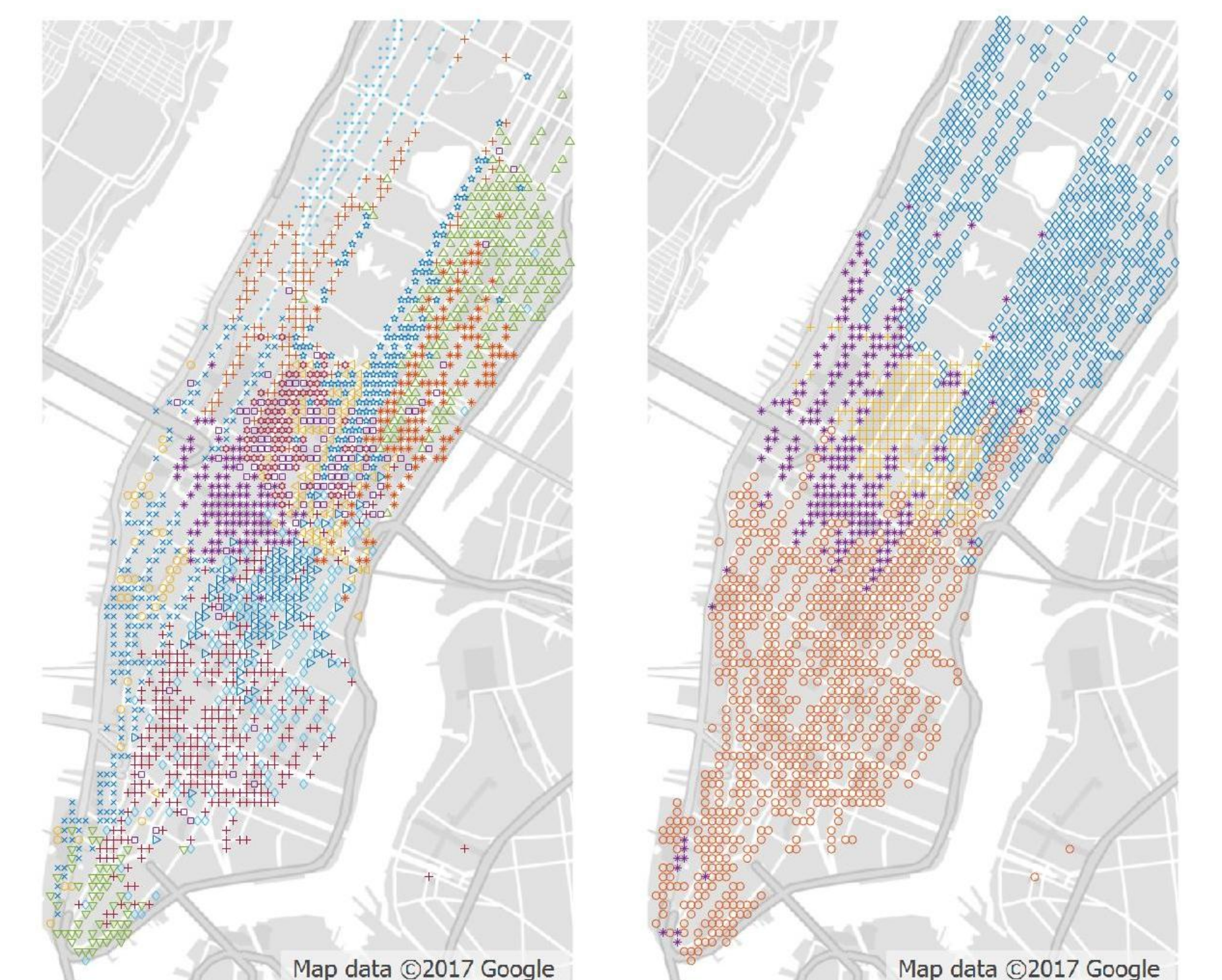


We discretize the map into a fine grid and model each taxi trip as a single state transition sample generated by an implicit city-wide random walk.



GHA: 4 clusters

GHA: 10 clusters



GHA: 15 clusters

SVD: 4 clusters

A practically impressive observation is: our algorithm uses less than 1 Mbytes memory for $r = 4, 10, 15$. In contrast, the batch partition uses about 200 Mbytes memory even for $r = 4$.