

# A System of Experimental Design

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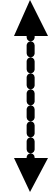
- System has four broad branches :
  - (i) regular orthogonal designs,
  - (ii) nonregular orthogonal designs,
  - (iii) response surface designs,
  - (iv) optimal designs.
- New opportunities in boarder areas.
  - Interface between (ii)  $\Leftrightarrow$  (iii), (iii)  $\Leftrightarrow$  (iv).
  - Space-filling designs.

New materials available in

“Experiments: Planning, Analysis and Parameter Design Optimization” by Wu - Hamada (2000)

Regular orthogonal designs ( Fisher, Yates, Finney, ...):  $2^{n-k}$ ,  $3^{n-k}$  designs, using minimum aberration criterion

Nonregular orthogonal designs (Plackett-Burman, Rao, Bose): Plackett-Burman designs, orthogonal arrays



(factor screening, projection)

Response surface designs (Box) : fitting a parametric response surface

Optimal designs (Kiefer): optimality driven by specific model/criterion

# Fundamental Principles for Factorial Effects

- **Effect Hierarchy Principle:**
  - Lower order effects more important than higher order effects
  - Effects of same order equally important
- **Effect Sparsity Principle:** Number of relatively important effects is small
- **Effect Heredity Principle:** for an interaction to be significant, at least one of its parent factors should be significant

# Fractional Factorial Designs

Run	1	2	3	12	13	23	123
1	-	-	-	+	+	+	-
2	-	-	+	+	-	-	+
3	-	+	-	-	+	-	+
4	-	+	+	-	-	+	-
5	+	-	-	-	-	+	+
6	+	-	+	-	+	-	-
7	+	+	-	+	-	-	-
8	+	+	+	+	+	+	+

col “12” = (col 1)  $\times$  (col2), etc.

- 4 factors: 1, 2, 3, 4 = 12  
(4 & 12 are said to be **aliased**)  
 $2^{4-1}$  design: I = 124
- 5 factors: 1, 2, 3, 4 = 12, 5 = 13  
 $2^{5-2}$  design: I = 124 = 135 = 2345  
(defining contrast subgroup)

- Resolution = shortest wordlength in the defining contrast subgroup of a design
- Design of same resolution can be quite different

$$d_1 : \quad \mathbf{I=4567=12346=12357}$$

$$d_2 : \quad \mathbf{I=1236=1457=234567}$$

both are  $2^{7-2}$  design of resolution IV

but  $d_1$  is better (why?)

- Let  $A_i(\mathbf{d})$  = no. of words of length  $i$  in the defining contrast subgroup of design  $\mathbf{d}$
- **Minimum aberration criterion** (Fries-Hunter, 1980): sequentially minimizes the values  $A_3, A_4, A_5, \dots$  etc.
- Aberration criterion is an extension of resolution criterion
- Ready-to-use tables of minimum aberration (and related)  $2^{k-p}$  designs in WH

# Thirty-Two Run Fractional Factorial Designs

k	F&R	Design Generators	Clear Effects
6	$2_{VI}^{6-1}$	6 = 12345	all six main effects, all 15 2fi's
7	$2_{IV}^{7-2}$	6 = 123, 7 = 1245	all seven main effects, 14, 15, 17, 24, 25, 27, 34, 35, 37, 45, 46, 47, 56, 57, 67
8	$2_{IV}^{8-3}$	6 = 123, 7 = 124, 8 = 1345	all eight main effects, 15, 18, 25, 28, 35, 38, 45, 48, 56, 57, 58, 68, 78
9	$2_{IV}^{9-4}$	6 = 123, 7 = 124, 8 = 125, 9 = 1345	all nine main effects, 19, 29, 39, 49, 59, 69, 79, 89
9	$2_{IV}^{9-4}$	6 = 123, 7 = 124, 8 = 134, 9 = 2345	all nine main effects, 15, 19, 25, 29, 35, 39, 45, 49, 56, 57, 58, 59, 69, 79, 89
10	$2_{IV}^{10-5}$	6 = 123, 7 = 124, 8 = 125, 9 = 1345, $t_0 = 2345$	all 10 main effects
10	$2_{III}^{10-5}$	6 = 12, 7 = 134, 8 = 135, 9 = 145, $t_0 = 345$	3, 4, 5, 7, 8, 9, $t_0$ , 23, 24, 25, 27, 28, 29, $2t_0$ , 36, 46, 56, 67, 68, 69, $6t_0$

# Extensions of Minimum Aberration to Designs with Factor Asymmetry

- $2^{n-k}$  designs in  $2^q$  blocks: “treatment defining contrast subgroup”, and “block defining contrast subgroup” are intertwined
- $2^{n-k}$  parameter designs, control and noise factors (control-by-noise interaction key to robustness): modified effect hierarchy principle:  
 $c, n, cn$  (1<sup>st</sup> group),  $cc, ccn, nn$  (2<sup>nd</sup> group), etc.
- $2^{n-k}$  split-plot designs,  $n_1$  whole-plot factors,  $n_2$  split-plot factors:  $wp$  effects,  $wp \times sp$  effects,  $sp$  effects treated differently, two variance components

## Two Types of Fractional Factorial Designs:

- Regular ( $2^{n-k}$ ,  $3^{n-k}$  designs):  
columns of the design matrix form a group over a finite field; the interaction between any two columns is among the columns  
 $\Rightarrow$  any two factorial effects are either  
*orthogonal* or *fully aliased*
- Nonregular (mixed-level designs, orthogonal arrays)  
some pairs of factorial effects can be *partially aliased*  
 $\Rightarrow$  more complex aliasing pattern

# Orthogonal Arrays

- Two columns of a design matrix are orthogonal if all possible level combinations of the two columns appear equally often in the matrix
- An **orthogonal array**  $OA(N, s_1^{m_1} \dots s_k^{m_k})$  of strength two is an  $N \times m$  matrix,  
 $m = m_1 + \dots + m_k$  in which  $m_i$  columns have  $s_i$  levels and any two columns are orthogonal
- $2^{n-k}$ ,  $3^{n-k}$  designs are OA's

# Cast Fatigue Experiment

Factor	Level	
	-	+
<i>A.</i> initial structure	as received	$\beta$ treat
<i>B.</i> bead size	small	large
<i>C.</i> pressure treat	none	HIP
<i>D.</i> heat treat	anneal	solution treat/age
<i>E.</i> cooling rate	slow	rapid
<i>F.</i> polish	chemical	mechanical
<i>G.</i> final treat	none	peen



# Partial and Complex Aliasing

- For the 12-run Plackett-Burman design OA(12, 2<sup>11</sup>)

$$E\hat{\mathbf{b}}_i = \mathbf{b}_i + \frac{1}{3} \sum_{j,k \neq i} \pm \mathbf{b}_{jk}$$

**partial aliasing:** coefficient  $\pm \frac{1}{3}$

**complex aliasing:** 45 (=  $\binom{10}{2}$ ) partial aliases

- Traditionally complex aliasing was considered to be a disadvantage
- Standard texts pay little attention to this type of designs

# Useful Orthogonal Arrays

- Collection in WH

$$\begin{array}{lll} OA(12, 2^{11}) & OA(12, 3^1 2^4) & OA(18, 2^1 3^7) \\ OA(18, 6^1 2^6) & OA(20, 2^{19}) & OA(24, 3^1 2^{16}) \\ OA(24, 6^1 2^{14}) & OA(36, 2^{11} 3^{12}) & OA(36, 3^7 6^3) \\ OA(36, 2^8 6^3) & OA(48, 2^{11} 4^{12}) & OA(50, 2^1 5^{11}) \\ OA(54, 2^1 3^{25}) & & \end{array}$$

- Run Size Economy

$$OA(12, 2^{11}) \text{ vs. } 16\text{-run } 2^{k-p} \text{ designs, } 8 \leq k \leq 11$$

$$OA(18, 3^7) \text{ vs. } 27\text{-run } 3^{k-p} \text{ designs, } 5 \leq k \leq 7$$

- Flexibility in level combinations

# Blood Glucose Experiment

$OA(18.2^13^7)$

Run	Factor								Mean Reading
	A	G	B	C	D	E	F	H	
1	0	0	0	0	0	0	0	0	97.94
2	0	0	1	1	1	1	1	1	83.40
3	0	0	2	2	2	2	2	2	95.88
4	0	1	0	0	1	1	2	2	88.86
5	0	1	1	1	2	2	0	0	106.58
6	0	1	2	2	0	0	1	1	89.57
7	0	2	0	1	0	2	1	2	91.98
8	0	2	1	2	1	0	2	0	98.41
9	0	2	2	0	2	1	0	1	87.56
10	1	0	0	2	2	1	1	0	88.11
11	1	0	1	0	0	2	2	1	83.81
12	1	0	2	1	1	0	0	2	98.27
13	1	1	0	1	2	0	2	1	115.52
14	1	1	1	2	0	1	0	2	94.89
15	1	1	2	0	1	2	1	0	94.70
16	1	2	0	2	1	2	0	1	121.62
17	1	2	1	0	2	0	1	2	93.86
18	1	2	2	1	0	1	2	0	96.10

# Analysis Strategies

- Traditionally experiments with **complex aliasing** were used for **screening** purpose, i.e., estimating main effects only
- A paradigm shift: using effect sparsity/heredity, Hamada-Wu (1992) recognized that complex aliasing can be turned into an advantage for studying interactions
- Analysis methods (frequentist and Bayesian) allow two-factor interactions to be entertained (in addition to main effects). Effective if the number of significant interactions is small

# Examples

- Cast Fatigue Experiment:

Main effect analysis: F (R<sup>2</sup>=0.45)

F, D (R<sup>2</sup>=0.59)

HW analysis: F, FG (R<sup>2</sup>=0.89)

F, FG, D (R<sup>2</sup>=0.92)

- Blood Glucose Experiment:

Main effect analysis: E<sub>q</sub>, F<sub>q</sub> (R<sup>2</sup>=0.36)

HW analysis: B<sub>1</sub>, (BH)<sub>1q</sub>, (BH)<sub>qq</sub> (R<sup>2</sup>=0.89)

Bayesian analysis also identifies B<sub>1</sub>, (BH)<sub>1l</sub>, (BH)<sub>1q</sub>, (BH)<sub>qq</sub> as having the highest posterior model probability

# Further Analysis

- Success in the HW analysis strategy led to research on the *hidden projection* properties of nonregular designs. Commonly used arrays like  $OA(12, 2^{11})$ ,  $OA(18, 3^7)$ ,  $OA(36, 2^{11}3^{12})$  have desirable projection properties (i.e., for 4 - 8 factors, a number of interactions can be estimated with good efficiency)
  - This is achieved *without* adding new runs
  - It has also inspired a new approach to response surface methodology
- Further Analysis

# A Poorman's Response Surface Methodology

- Consider an experiment to study three quantitative factors with up to 5 levels.

Factors and Levels, Ranitidine Experiment

Factor	Levels
A. pH	2, 3.42, 5.5, 7.58, 9
B. voltage (kV)	9.9, 14, 20, 26, 30.1
C. $\alpha$ -CD (mM)	0, 2, 5, 8, 10

- The design matrix and the data are given on the next page. The design differs from  $2^{k-p}$  design in two respects :
  - 6 replicates at the center,
  - 6 runs along the three axes.

It belongs to the class of *central composite designs*.

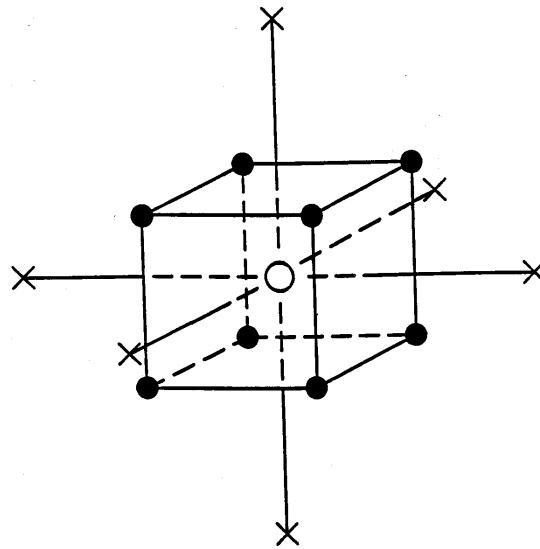
# Ranitidine Experiment

## Design Matrix and Response Data

Run	Factor			CEF	ln CEF
	A	B	C		
1	-1	-1	-1	17.293	2.850
2	1	-1	-1	45.488	3.817
3	-1	1	-1	10.311	2.333
4	1	1	-1	11757.084	9.372
5	-1	-1	1	16.942	2.830
6	1	-1	1	25.400	3.235
7	-1	1	1	31697.199	10.364
8	1	1	1	12039.201	9.396
9	0	0	-1.67	7.474	2.011
10	0	0	1.67	6.312	1.842
11	0	-1.68	0	11.145	2.411
12	0	1.68	0	6.664	1.897
13	-1.68	0	0	16548.749	9.714
14	1.68	0	0	26351.811	10.179
15	0	0	0	9.854	2.288
16	0	0	0	9.606	2.262
17	0	0	0	8.863	2.182
18	0	0	0	8.783	2.173
19	0	0	0	8.013	2.081
20	0	0	0	8.059	2.087

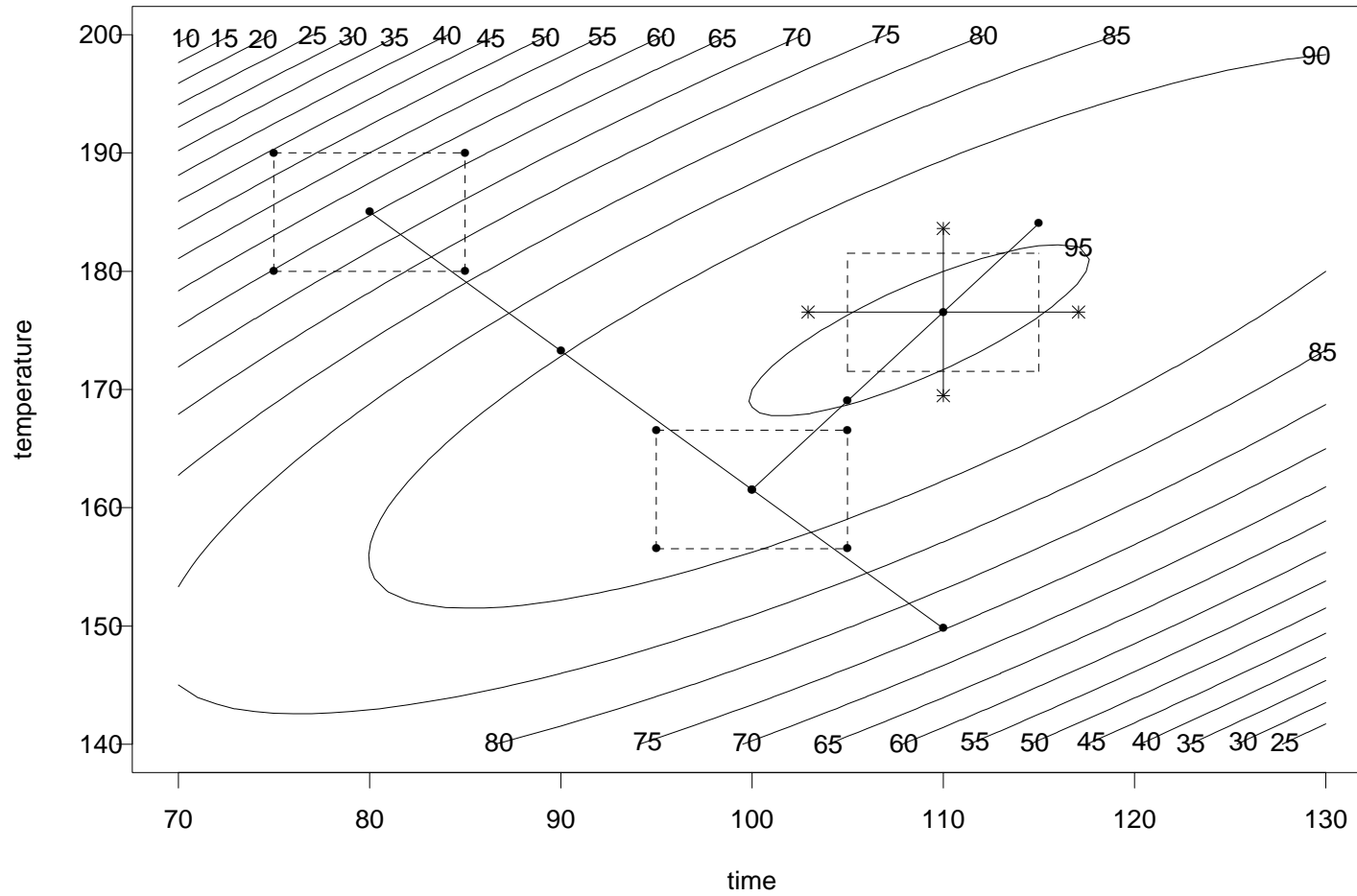
# Central Composite Designs

A simple CCD is shown graphically below. It has three parts  
(1) *cube* ( or corner) points, (2) *axial* (or star) points,  
(3) *center* points.



A Central Composite Design in Three Dimensions (cube point (dot), star point(cross), center point (circle))

# Sequential Exploration of Response Surface



# An Alternative to Standard Response Surface Methodology

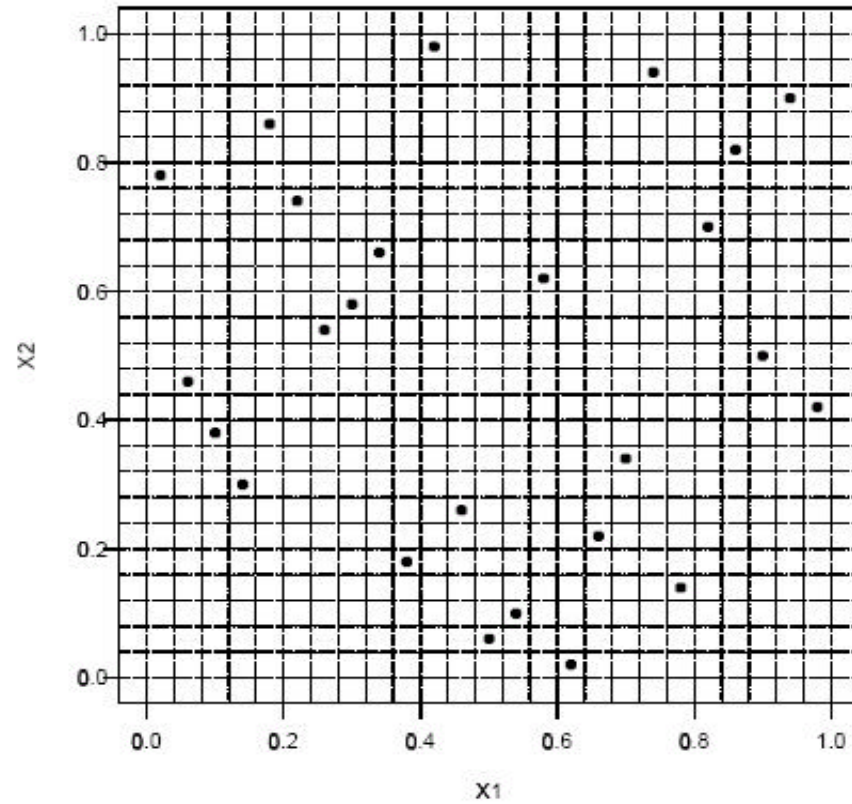
- Standard RSM employs a 2-stage experimentation strategy; this can be time consuming and expensive.
- S. W. Cheng - Wu (2001) proposed a new strategy to perform factor screening (1<sup>st</sup> order model) and response surface exploration (2<sup>nd</sup> order model) on the *same* experiment using *one* design, based on new optimality criteria called *projection - aberration*

# Optimal Designs

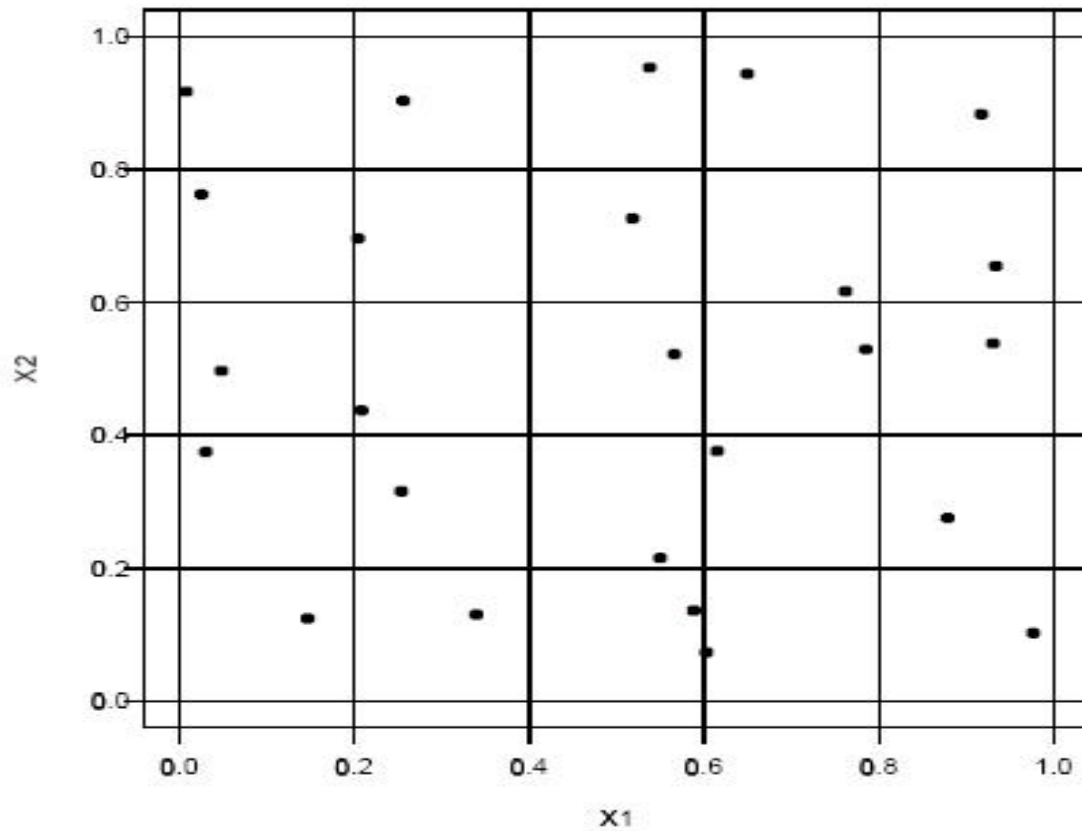
- $D$ -,  $G$ -,  $I$ -optimality based on a *single* model, performance not guaranteed over a variety of models. Performance is highly *model-dependent*.
- Exact optimality more interesting than approximate (continuous) optimality. Algorithms make more impact than theory; generally *applicable to any models*.
- Bigger impact when used in conjunction with or as a supplement to a combinatorial or reasonably uniform design (irregular design, follow-up experiment, sequential designs).
- Generally useful as a benchmark.

# Space-Filling Designs

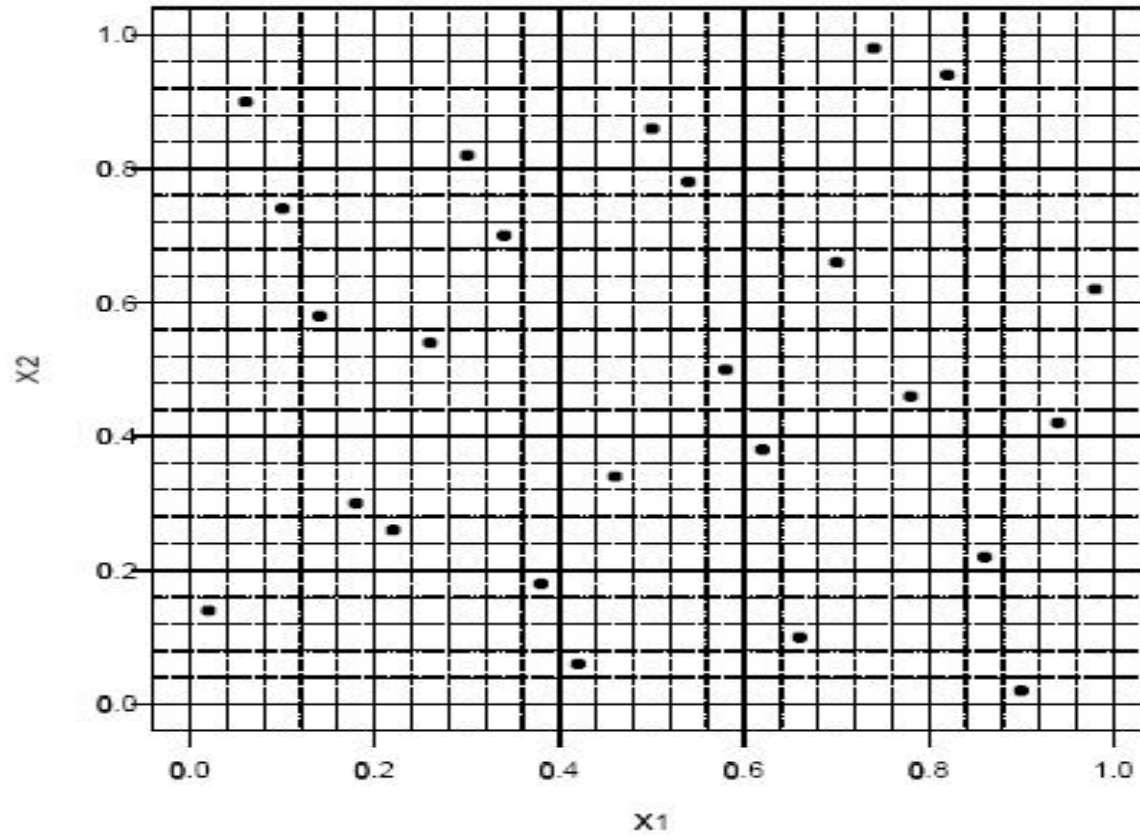
- Latin Hypercube in  $s^d$  (McKay, Beckman, Conover, 1979): *one-dimensional balance* for each of the  $d$  dimensions with  $s$  levels. Various extensions available: OA-based LHS (Tang, 1993) achieves 2- or higher-dimensional balance by combining LHS and OA.
- Uniform designs (Fang): based on wrapping around a sequence in high dimensions using number-theoretic justifications. More generally, *good lattice points*: work by Niederreiter etc.
- Design criteria and modeling are drastically *different* from all the previous approaches: space-filling, minimax or maximin distance; use of semi-parametric modeling or Gaussian process modeling that allows highly *nonlinear* fitting.



25 points of a Latin hypercube sample



25 points of a randomly centered randomized orthogonal array. For any two variables, there is one point in each reference square.



25 points of an OA-based Latin hypercube sample

# Innovations in Bayesian Analysis for Designed Experiments

- Choice of priors reflects the three principles (hierarchy, sparsity, heredity)
- Model search strategy depends on nature of design (much easier for regular designs; not so for nonregular designs); strategy should exploit the effect aliasing pattern
- Convenient for computer experiments.

# A Summary of Layout Techniques

A flexible strategy for selecting a suitable design from among

- (i)  $2^{k-p}$  designs : 8, 16, 32, 64, 128-run designs.
- (ii)  $3^{k-p}$  designs : 9, 27, 81-run designs (optimal selection of factor columns by minimum aberration criterion).
- (iii) Plackett-Burman designs :  $OA(12, 2^{11}), OA(20, 2^{19}), \dots$
- (iv) Mixed-level designs :  $OA(8, 4^1 2^4), OA(16, 4^m 2^n)$  (derived from  $2^{k-p}$  designs),  $OA(18, 2^1 3^7), OA(36, 2^{11} 3^{12}), OA(12, 3^1 2^4), OA(24, 3^1 2^{16}), \dots$   
(allowing main effects and a flexible choice of interactions to be estimated; hidden-projection property).
- (v) Central Composite designs.
- (vi) Optimal Designs (SAS/QC).
- (vii) Space-filling designs (Latin hypercube designs).