

PROBLEM-1: SOLUTION SUBMITTED BY GITA AYU GIVEN BELOW.

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 HW #6

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Problem 1

3^{5-2} design with $D = ABC$ and $E = ABC^2$

a. Defining contrast subgroup: $I = ABCD^2 = ABC^2E^2 = ABDE = CDE$

The 3^{5-2} design has 27 runs which has total of 13 sets of aliased effects, shown below:

I	=	ABCD ²	=	ABC ² E ²	=	ABDE	=	CDE ²		
A ²	=	A	=	BCD ²	=	BC ² E ²	=	BDE	=	AC ² D ² E
			=	AB ² C ² D	=	AB ² CE	=	AB ² D ² E ²	=	ACDE ²
B ²	=	B	=	ACD ²	=	AC ² E ²	=	ADE	=	BC ² D ² E
			=	AB ² CD ²	=	AB ² C ² E ²	=	AB ² DE	=	BCDE ²
C ²	=	C	=	ABD ²	=	ABCE ²	=	ABC ² DE	=	DE ²
			=	ABC ² D ²	=	ABE ²	=	ABCDE	=	CD ² E
D ²	=	D	=	ABCD	=	ABC ² D ² E ²	=	ABE	=	CE ²
			=	ABC	=	ABC ² DE ²	=	ABD ² E	=	CD ² E ²
E ²	=	E	=	ABCD ² E ²	=	ABC ² E	=	ABD	=	CDE
			=	ABCD ² E	=	ABC ²	=	ABDE ²	=	CD
A ² B ²	=	AB	=	CD ²	=	CE	=	DE	=	ABC ² D ² E
			=	ABC ² D	=	ABCE	=	ABD ² E ²	=	ABCDE ²
A ² B	=	AB ²	=	BC ² D	=	BCE	=	BD ² E ²	=	AB ² C ² D ² E
			=	AC ² D	=	ACE	=	AD ² E ²	=	AB ² CDE ²
A ² C ²	=	AC	=	BD ²	=	BCE ²	=	BC ² DE	=	AD ² E
			=	AB ² CD	=	AB ² E	=	AB ² C ² D ² E ²	=	AC ² DE ²
A ² C	=	AC ²	=	BC ² D ²	=	BE ²	=	BCDE	=	ACD ² E
			=	AB ² D	=	AB ² C ² E	=	AB ² CD ² E ²	=	ADE ²
A ² D ²	=	AD	=	BCD	=	BC ² D ² E ²	=	BE	=	AC ² E
			=	AB ² C ²	=	AB ² CD ² E	=	AB ² DE ²	=	ACD ² E ²
A ² D	=	AD ²	=	BC	=	BC ² DE ²	=	BD ² E	=	C ² DE
			=	AB ² C ² D ²	=	AB ² CDE	=	AB ² E ²	=	ACE ²
A ² E ²	=	AE	=	BCD ² E ²	=	BC ² E	=	BD	=	AC ² D ² E ²
			=	AB ² C ² DE ²	=	AB ² C	=	AB ² D ² E	=	ACD
A ² E	=	AE ²	=	BCD ² E	=	BC ²	=	BDE ²	=	AC ² D ²
			=	AB ² C ² DE	=	AB ² CE ²	=	AB ² D ²	=	ACDE

b. Clear main effects are: A, B

Clear two-factor interactions are: AB²

c. Word length pattern is, $W = (1,3,0)$

Compare the 3^{5-2} design above ($D = ABC$ and $E = ABC^2$) with the minimum aberration design of 3^{5-2} design with $D = AB$ and $E = AB^2C$ available on page 250. From page 216, the defining relations and the number of words length i in the subgroup are calculated. By using the information available on page 216, we obtained that the current design and the minimum aberration design have the same word length of $W = (1,3,0)$. Because they are identical, we can conclude that the design above is another minimum aberration design besides 3^{5-2} design with $D = AB$ and $E = AB^2C$.

PROBLEM-2

Clearly, $B=CE$, $D=ACE$ and $F=A2CE$.

Thus, the defining contrast subgroup is

$$I = BC^2E^2 = ACD^2E = AC^2E^2F = ABD^2 = AB^2C^2D^2E^2 = ABC^2EF = AB^2F = ADF^2 = CDEF = AB^2CDEF^2 = BCD^2EF^2 = BDF = ABC^2DE^2F^2.$$

It can easily be seen that $AC = DE^2$, $AC^2 = EF^2$, $AE = CD^2$, $AE^2 = CF^2$ and thus none of them are clear.

The aliasing relations of main effects and two-factor interactions are:

$$\begin{aligned} A &= BD^2 = BF^2 = DF^2 \\ B &= AD^2 = AF = CE = DF \\ C &= BE^2 \\ D &= AB = AF^2 = BF \\ E &= BC^2 \\ F &= AB^2 = AD = BD \\ AC &= DE^2 \\ AC^2 &= EF^2 \\ AE &= CD^2 \\ AE^2 &= CF^2 \\ CD &= EF \\ CE^2 &= BC = BE \\ CF &= DE \end{aligned}$$

Thus, in the ANOVA, besides $A, B, C, D, E, F, AC, AC^2, AE, AE^2$, one can choose one from (CD, EF) , one from (CE^2, BC, BE) and one from (CF, DE)

METHOD-1: SIMPLE ANALYSIS

ANALYSIS OF VARIANCE TABLE

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
factor(A)	2	47.620	23.810	2.8297	0.0766119 .
factor(B)	2	11.822	5.911	0.7025	0.5041863
factor(C)	2	39.586	19.793	2.3523	0.1143593
factor(D)	2	188.037	94.019	11.1736	0.0002913 ***
factor(E)	2	80.433	40.216	4.7795	0.0167104 *
factor(F)	2	9.933	4.966	0.5902	0.5611903
factor(AC)	2	17.405	8.703	1.0343	0.3691520
factor(AC2)	2	54.697	27.348	3.2502	0.0543543 .
factor(AE)	2	232.711	116.355	13.8282	7.333e-05 ***
factor(AE2)	2	3.898	1.949	0.2316	0.7947942
factor(CD)	2	61.454	30.727	3.6517	0.0394778 *
factor(BC)	2	4.238	2.119	0.2519	0.7791619
factor(DE)	2	24.141	12.071	1.4345	0.2558124
Residuals	27	227.187	8.414		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

The significant effects (at 5% level) are D, E, AE and CD. From the interaction plots, we select the best levels as $D=2$, $E=0$, $A=0$, $C=2$, $D=2$.

METHOD-2: LINER AND QUADRATIC CONTRASTS

FINAL OUTPUT OF STEPWISE REGRESSION WITH ALL LINEAR AND QUADRATIC EFFECTS

Call:

```
lm(formula = y ~ Al + Cl + Dl + El + Dq + Eq + AC1q + ACq1 + ACqq + AE1q + CD11 + CF1q, data = design.sel)
```

Coefficients:

(Intercept)	Al	Cl	Dl	El	Dq
12.157	1.552	-1.424	2.082	-1.247	2.472
Eq	AC1q	ACq1	ACqq	AE1q	CD11
1.707	2.045	-2.213	1.361	-3.507	5.023
CF1q					
2.391					

The above model is too large; and we should see whether all terms are significant by displaying the t-statistics.

```
> g=lm(y~Al+Cl+Dl+El+Dq+Eq+AC1q+ACq1+ACqq+AE1q+CD11+CF1q)
> summary(g)
```

Call:

```
lm(formula = y ~ Al + Cl + Dl + El + Dq + Eq + AC1q + ACq1 + ACqq + AE1q + CD11 + CF1q)
```

Residuals:

Min	1Q	Median	3Q	Max
-4.31906	-1.66155	0.06925	1.86304	5.44715

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	12.1574	0.3610	33.676	< 2e-16 ***
Al	1.5517	0.6253	2.482	0.017271 *
Cl	-1.4240	0.6253	-2.277	0.028048 *
Dl	2.0820	0.6253	3.330	0.001846 **
El	-1.2473	0.6253	-1.995	0.052756 .
Dq	2.4722	0.6253	3.954	0.000297 ***
Eq	1.7067	0.6253	2.729	0.009304 **
AC1q	2.0448	1.0830	1.888	0.066117 .
ACq1	-2.2132	1.0830	-2.043	0.047467 *
ACqq	1.3611	1.0830	1.257	0.215958
AE1q	-3.5069	1.4441	-2.428	0.019635 *
CD11	5.0231	1.2506	4.017	0.000246 ***
CF1q	2.3909	1.3017	1.837	0.073491 .

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.653 on 41 degrees of freedom

Multiple R-Squared: 0.7124, Adjusted R-squared: 0.6282

F-statistic: 8.461 on 12 and 41 DF, p-value: 9.639e-08

ITERATION 1

Thus we delete the terms that are not significant at 5% level (El, AC1q, ACqq, CF1q) and re-run the regression.

```
> g1=lm(y~Al+Cl+Dl+Dq+Eq+ACq1+AE1q+CD11)
```

```
> summary(g1)
```

Call:

```
lm(formula = y ~ Al + Cl + Dl + Dq + Eq + ACq1 + AE1q + CD11)
```

Residuals:

Min	1Q	Median	3Q	Max
-5.14441	-1.77221	-0.02101	1.65487	5.91435

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	12.1574	0.3936	30.886	< 2e-16	***
Al	1.5517	0.6818	2.276	0.02766	*
Cl	-1.4240	0.6818	-2.089	0.04242	*
Dl	2.0820	0.6818	3.054	0.00379	**
Dq	2.4722	0.6818	3.626	0.00073	***
Eq	1.7067	0.6818	2.503	0.01600	*
ACq1	-2.2132	1.1809	-1.874	0.06741	.
AElq	-2.0355	1.3101	-1.554	0.12725	.
CD11	5.6603	1.3101	4.321	8.48e-05	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.893 on 45 degrees of freedom
Multiple R-Squared: 0.6247, Adjusted R-squared: 0.558
F-statistic: 9.362 on 8 and 45 DF, p-value: 1.711e-07

ITERATION 2

We again delete the terms that are not significant at 5% level (ACq1, Aelq) and re-run the regression.

```
> g2=lm(y~Al+Cl+Dl+Dq+Eq+CD11)
> summary(g2)
```

```
Call:
lm(formula = y ~ Al + Cl + Dl + Dq + Eq + CD11)
```

```
Residuals:
    Min     1Q   Median     3Q     Max
-5.981 -1.395 -0.228  2.131  5.914
```

```
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 12.1574     0.4097  29.671 < 2e-16 ***
Al           1.5517     0.7097   2.186 0.03379 *
Cl          -1.4240     0.7097  -2.007 0.05057 .
Dl           2.0820     0.7097   2.934 0.00516 **
Dq           2.4722     0.7097   3.483 0.00108 **
Eq           1.7067     0.7097   2.405 0.02017 *
CD11         6.5417     1.2292   5.322 2.81e-06 ***
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.011 on 47 degrees of freedom
Multiple R-Squared: 0.5753, Adjusted R-squared: 0.521
F-statistic: 10.61 on 6 and 47 DF, p-value: 2.003e-07

ITERATION 3

We again delete the term that are not significant at 5% level (Cl) and re-run the regression.

```
> g3=lm(y~Al+Dl+Dq+Eq+CD11)
> summary(g3)
```

```
Call:
lm(formula = y ~ Al + Dl + Dq + Eq + CD11)
```

```
Residuals:
    Min     1Q   Median     3Q     Max
-5.73843 -2.01620 -0.09954  2.36921  6.83796
```

```
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 12.1574     0.4097  29.671 < 2e-16 ***
Al           1.5517     0.7097   2.186 0.03379 *
Dl           2.0820     0.7097   2.934 0.00516 **
Dq           2.4722     0.7097   3.483 0.00108 **
Eq           1.7067     0.7097   2.405 0.02017 *
CD11         6.5417     1.2292   5.322 2.81e-06 ***
```

(Intercept)	12.1574	0.4225	28.778	< 2e-16	***
Al	1.5517	0.7317	2.121	0.03915	*
Dl	2.0820	0.7317	2.845	0.00650	**
Dq	2.4722	0.7317	3.379	0.00145	**
Eq	1.7067	0.7317	2.332	0.02392	*
CD11	6.5417	1.2674	5.162	4.64e-06	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.104 on 48 degrees of freedom
Multiple R-Squared: 0.5389, Adjusted R-squared: 0.4908
F-statistic: 11.22 on 5 and 48 DF, p-value: 3.395e-07

Thus, the final model is

$$Y = 1.5517 Al + 2.0820 Dl + 2.4722 Dq + 1.7067 Eq + 6.5417 CD11$$

To maximize y , we choose $Al = +1$, $Dl = +1$, $Dq = +1$, $Eq = +1$ and $CD11 = +1$ (which means either $C1=D1=-1$ or $C1=D1=1$).

Finally, we choose the optimum combination as follows:

- i. $Al = +1$ means $A = 2$.
- ii. $Dl = +1$ and $Dq = +1$ means $D = 2$.
- iii. $Eq = +1$ means $E = 0$ or $E = 2$; looking at the main effects plot we choose $E = 0$.
- iv. Since $Dl = 1$, we choose $C1 = 1$ i.e., $C = 2$.

Disagreement between results obtained from method-1 and method-2:

The only difference between the settings obtained from method-1 and method-2 is the setting of factor A, which is 0 in method 1 and 2 in method 2. This apparent anomaly can be explained very easily. Note that in the simple analysis, the setting of A was obtained on the basis of the $A \times E$ interaction plot. A careful observation of this plot tells us that choosing $A = 0$ or $A = 2$ makes very little difference. However, if we have the main effect of A in the model, one would go for $A = 2$. Thus, considering both the main effect of A and interaction AE , $A = 2$ may be a slightly better choice.